



## METHOD

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### Key Points:

- Model reproduces O<sub>2</sub>-CO<sub>2</sub> dynamics in stream water
- Stream CO<sub>2</sub> outgassing is partitioned between internal and external sources
- Estimated O-C stoichiometry is different from unity

### Supporting Information:

Supporting Information may be found in the online version of this article.

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# A Coupled O<sub>2</sub>-CO<sub>2</sub> Model for Joint Estimation of Stream Metabolism, O-C Stoichiometry, and Inorganic Carbon Fluxes

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**Abstract** We determine where stream carbon dioxide (CO<sub>2</sub>) comes from by developing a model for the joint estimation of stream metabolism, oxygen-carbon (O-C) stoichiometry, and fluxes of dissolved inorganic carbon (DIC), based on observations of stream oxygen (O<sub>2</sub>) and CO<sub>2</sub> concentrations. The model is based on a stream reach mass balance of O<sub>2</sub>, DIC, and total alkalinity, and it accounts for the carbonate system and the contribution of lateral flow. O<sub>2</sub> and DIC mass balances are coupled through stoichiometric coefficients for photosynthesis and combined autotrophic and heterotrophic respiration. Under the assumption of constant alkalinity and circumneutral pH, the model simplifies and includes 8 parameters, which are estimated through a Bayesian hierarchical framework. The model accurately reproduced time series of O<sub>2</sub> and CO<sub>2</sub> from three diverse sites across size and carbonate chemistry gradients. Results allow partitioning of the stream DIC budget, and thus the source of stream CO<sub>2</sub> outgassing, into internal (in-stream net ecosystem production) and external (lateral input of terrestrial DIC and atmospheric input) contributions. We observed that the estimated stoichiometric coefficients were typically different from 1—contrary to typical assumptions—leading to divergent estimates of stream CO<sub>2</sub> sources depending on the measurement (i.e., O<sub>2</sub> vs. C). Parameter posterior distributions revealed the source of parameter uncertainty and the equifinality of some processes in reproducing stream CO<sub>2</sub> dynamics, suggesting targeted variables to further investigate in order to better constrain stream C balance. The proposed model is a useful tool for incorporating the rapidly growing stream CO<sub>2</sub> data sets into our understanding of terrestrial-aquatic C linkages.

**Plain Language Summary** The metabolism of streams and rivers is defined by two key processes: production, during which autotrophic organisms synthesize organic carbon by absorbing carbon dioxide (CO<sub>2</sub>) and releasing oxygen (O<sub>2</sub>), and respiration, where organic carbon is broken down, consuming O<sub>2</sub> and releasing CO<sub>2</sub> and energy. Because production occurs only during daylight hours, when light is available for photosynthesis, this leads to diel fluctuations in the concentrations of O<sub>2</sub> and CO<sub>2</sub> in stream water. In this study, we developed a model that leverages joint observations of O<sub>2</sub> and CO<sub>2</sub> to estimate not only stream metabolism but also the ratio of carbon to oxygen usage and the lateral fluxes of inorganic carbon leaking from terrestrial to stream ecosystems. The results address two emerging issues in freshwater biogeosciences: accurately estimating CO<sub>2</sub> emissions and closing the carbon budget at catchment, regional, and global scales.

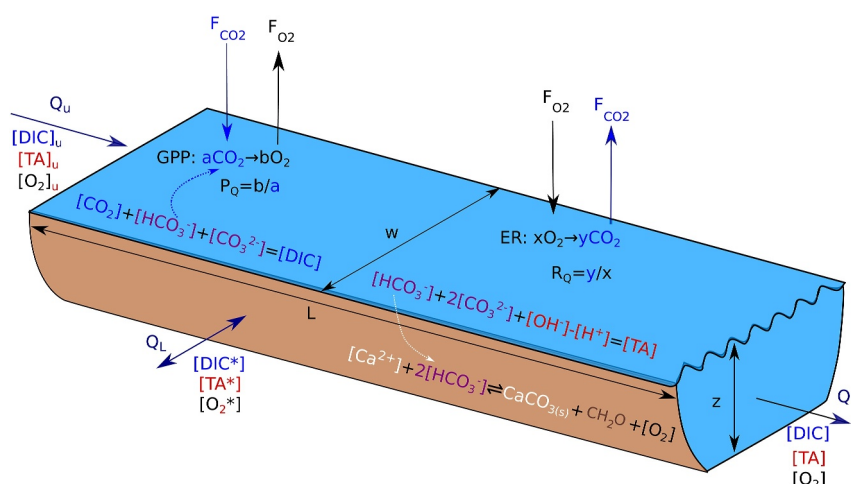
## 1. Introduction

There is an unexplained gap (ca. 1.5 Pg-C y<sup>-1</sup>) between estimates of terrestrial carbon (C) lateral export to receiving waters and resulting freshwater CO<sub>2</sub> emissions (Battin et al., 2009; Kirk & Cohen, 2023; Regnier et al., 2013). The ultimate source of observed freshwater carbon dioxide (CO<sub>2</sub>) emissions therefore remains an important uncertainty in our understanding of the connection between terrestrial and freshwater ecosystems and ultimately the global C cycle (Battin et al., 2023; Bernhardt et al., 2022). Streams and rivers represent the largest share of freshwater CO<sub>2</sub> emissions (Drake et al., 2018), driving research focus to their dynamics and raising the question of whether, and in what proportion, their observed CO<sub>2</sub> is (a) produced from in-stream respiration of organic carbon or (b) delivered directly to streams from terrestrial soils as dissolved inorganic carbon (DIC). Resulting from this interest are an increasing number of subdaily resolution CO<sub>2</sub> data sets (<https://data.stream-pulse.org>, Appling, Read, et al. (2018)) that have the potential to reveal the relative magnitudes of so-called internal versus external C sources (Bernal et al., 2022; Hotchkiss et al., 2015; Kirk & Cohen, 2023).

A long history of metabolism estimates derived from oxygen (O<sub>2</sub>) concentration time series informs our understanding of stream ecosystem respiration (ER) and gross primary production (GPP) (Bernhardt et al., 2022;

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**Figure 1.** Conceptual setup of the model domain. State variables and their component chemical species are shown in like colors DIC, blue; TA, red; and  $O_2$ , black. Shared component species among TA and DIC are shown in purple. Also shown are relevant fluxes including GPP, ER (and their DIC- $O_2$  coupling coefficients), and biogenic calcite precipitation (white), which removes TA in addition to DIC. A blue dotted line from  $HCO_3^-$  to  $CO_2$  indicates that the full model does not distinguish between these species for GPP uptake of DIC, as autotrophic uptake of both  $HCO_3^-$  and  $CO_2$  induces identical changes in pH, and neither pathway affects TA.

Mulholland et al., 2001; Odum, 1956). Indeed, hundreds of stream reaches across nearly 70 years have employed a  $O_2$  mass balance method to simultaneously determine GPP, ER, and atmospheric gas exchange (Demars et al., 2015). By contrast, direct estimates of external  $CO_2$  transport to streams are far more limited as they require measurements of both lateral flow and  $CO_2$  concentration in groundwater (Kirk & Cohen, 2023). Thus, of the two potential stream  $CO_2$  sources, internal generation via heterotrophic net ecosystem production ( $NEP = GPP - ER$ ,  $ER > GPP$ ) is the most constrained (Hotchkiss et al., 2015), leading to estimation of external sources by difference. There is an untapped potential in stream  $CO_2$  time series to directly estimate both sources and therefore begin to close the C gap between terrestrial and freshwater ecosystems.

Inherent to the  $O_2$  balance method is an assumption of a fixed ratio (1:1 or sometimes 1.2:1) between  $O_2$  and C based on the production and consumption of organic carbon by photosynthesis and respiration, respectively (Aho et al., 2021a; Hotchkiss et al., 2015). Contemporary evidence suggests that this stoichiometry—referred to as the photosynthetic (mol  $O_2$ /mol  $CO_2$ ,  $P_Q$ ) and respiratory (mol  $CO_2$ /mol  $O_2$ ,  $R_Q$ ) quotients—varies widely depending on environmental conditions (Berggren et al., 2012; Trentman et al., 2023). This implies that our understanding of C fluxes based on decades of  $O_2$  modeling may be biased, potentially explaining a substantial proportion of variation in internal  $CO_2$  production. Thus, there is an obvious need to couple  $O_2$  and  $CO_2$  dynamics to constrain uncertainty in these key stoichiometric parameters.

We developed a reach-scale river model of subdaily  $CO_2$  and  $O_2$  dynamics based on mass balance of  $O_2$ , DIC, and total alkalinity, and used it to jointly estimate metabolic fluxes (GPP and ER), the stoichiometric quotients ( $P_Q$  and  $R_Q$ ), and the fluxes of DIC, and, as a result, the magnitude of external and internal sources of  $CO_2$  outgassing. We hypothesized that the model would not generate substantially different values for GPP, ER, or gas exchange compared to  $O_2$  mass balance approaches as these are relatively robust. However, we expected that the model would improve understanding of  $P_Q$ ,  $R_Q$ , and lateral DIC flux given that lateral discharge was constrained.

## 2. Methods

### 2.1. Model Equations

We consider a stream reach of length  $L$  [m], average width  $w$  [m], and time varying average depth  $z(t)$  [m], which stores a water volume  $V(t) = w z(t)L$  [ $m^3$ ]. The reach receives inflow from upstream,  $Q_u(t)$  [ $m^3 d^{-1}$ ], and laterally,  $Q_L(t)$ , and discharges streamflow to the downstream reach,  $Q(t)$  (Figure 1). Note that in the case of a losing stream,  $Q_L(t)$  is negative and represents a flow output.

The model equations are derived from the coupled reach-scale mass balance of water and three quantities that are conservative under mixing: oxygen ( $O_2$ ), dissolved inorganic carbon (DIC), and total alkalinity (TA). Note that mass balance of water translates into a volume balance when water is approximated as an incompressible fluid. The mass balance is driven by hydrologic flows, GPP and ER, gas exchange at the water-air interface ( $F$ ), possible calcite precipitation/dissolution, and advection and mixing in the stream water. The model assumes that the reach is a well-mixed reactor, and thus, the mass balances translate into the following set of ordinary differential equations:

$$\frac{dV(t)}{dt} = Q_u(t) + Q_L(t) - Q(t) \quad (1)$$

$$\frac{dM_{O_2}(t)}{dt} = Q_u(t)O_{2,u}(t) + Q_L(t)O_2^*(t) - Q(t)O_2(t) + L \cdot w [GPP(t) - ER(t) + F_{O_2}(t)] \quad (2)$$

$$\begin{aligned} \frac{dM_{DIC}(t)}{dt} = & Q_u(t)DIC_u(t) + Q_L(t)DIC^*(t) - Q(t)DIC(t) \\ & + L \cdot w \left[ -\frac{GPP(t)}{P_Q(t)} + R_Q(t)ER(t) + F_{CO_2}(t) + g_{DIC}(t) \right] \end{aligned} \quad (3)$$

$$\frac{dM_{TA}(t)}{dt} = Q_u(t)TA_u(t) + Q_L(t)TA^*(t) - Q(t)TA(t) + L \cdot w \cdot g_{TA}(t) \quad (4)$$

where  $M_{O_2}(t)$ ,  $M_{DIC}(t)$ , and  $M_{TA}(t)$  represent mass (in [mol] and [Eq] for TA) of the respective quantities, while  $O_2(t)$ ,  $DIC(t)$ , and  $TA(t)$  the corresponding well-mixed volumetric concentrations [ $\text{mol m}^{-3}$ ]. Concentrations with subscripts  $u$  refer to the concentrations in the upstream inflow. Concentrations with an asterisk (\*) are equal to the concentration of lateral flow when  $Q_L(t) > 0$ , and are equal to the stream water concentration when  $Q_L(t) < 0$  (losing conditions). The bracketed terms in the right-hand side of Equations 2–4 represent fluxes per unit of benthic stream area [ $\text{mol m}^{-2} \text{d}^{-1}$ ] that affect the state variable mass balance. They are  $GPP(t)$ ,  $ER(t)$ ,  $F_{O_2}(t)$ , and  $F_{CO_2}(t)$ . We note that because we account for the effect of GPP on DIC mass balance, and not on  $CO_2$ , this setup allows modeling the direct autotrophic  $HCO_3^-$  uptake, which is observed in  $CO_2$  limited ecosystems and in streams with high TA (Aho et al., 2021a; Maberly & Gontero, 2017). For completeness, we also indicate flux terms,  $g_{DIC}(t)$  and  $g_{TA}(t)$ , for additional biological and physical processes, such as calcification, which reduces both DIC and TA (Wolf-Gladrow et al., 2007), although these processes are not further considered in the remainder of the paper.  $P_Q$  and  $R_Q$  represent the photosynthetic (i.e., the molar ratio between  $O_2$  and DIC for primary production) and the respiratory (i.e., the molar ratio between  $CO_2$  and  $O_2$  for ecosystem respiration) quotients, which couple the  $O_2$  and DIC mass balances.

The model equations can be simplified by assuming that the upstream reach shares similar conditions with the focus reach, an inherent assumption in single-station  $O_2$  analyses (Hall & Hotchkiss, 2017). Specifically, this assumption implies a stream section upstream that (a) shares similar conditions with the focus reach and that (b) is sufficiently long to be unaffected by initial conditions (see the detailed derivation in Reichert et al., 2009). Using this assumption, the resulting model equations read (see detailed derivation in Supporting Information S1) as the following:

$$\frac{dO_2(t)}{dt} = \frac{GPP(t)}{z(t)} - \frac{ER(t)}{z(t)} + \frac{q_L(t)}{z(t)} [O_2^*(t) - O_2(t)] + K_{O_2}(t) [O_{2,sat}(t) - O_2(t)] \quad (5)$$

$$\begin{aligned} \frac{dDIC(t)}{dt} = & \frac{g_{DIC}(t)}{z(t)} - \frac{GPP(t)}{z(t)P_Q(t)} + R_Q \frac{ER(t)}{z(t)} + \frac{q_L(t)}{z(t)} [DIC^*(t) - DIC(t)] \\ & + K_{CO_2}(t) [CO_{2,sat}(t) - CO_2(t)] \end{aligned} \quad (6)$$

$$\frac{dTA(t)}{dt} = \frac{g_{TA}(t)}{z(t)} + \frac{q_L(t)}{z(t)} [TA^*(t) - TA(t)] \quad (7)$$

where  $q_L(t) = Q_L(t)/(w \cdot L)$  represents the lateral discharge per unit of streambed area [ $\text{m d}^{-1}$ ]. Stream-atmosphere gas exchange is modeled using Fick's Law with the atmospheric saturation concentration ( $\text{O}_{2,\text{sat}}$  and  $\text{CO}_{2,\text{sat}}$  [ $\text{mol m}^{-3}$ ]) and the gas exchange rates ( $K_{\text{O}_2}$  and  $K_{\text{CO}_2}$  [ $\text{d}^{-1}$ ]) (Wanninkhof et al., 2009). In the case of a losing reach, one has  $\text{DIC}^*(t) = \text{DIC}(t)$ ,  $\text{O}_2^*(t) = \text{O}_2(t)$ , and  $\text{TA}^*(t) = \text{TA}(t)$ , and thus, the terms related to  $q_L$  are null, which implies that the lateral flow has, as expected, no effect on these variables.

The exchange flux of DIC with the atmosphere is a function of  $\text{CO}_2$  concentration, invoking the need to calculate the water carbonate system. Equilibration of carbonate species occurs with a time scale (around minutes) shorter than the typical diel fluctuations of the quantities considered. Therefore, using the principle of timescale separation, we assume instantaneous equilibria among the carbonate species. At each time  $t$ , we calculate  $\text{CO}_2(t)$  [ $\text{mol m}^{-3}$ ] and  $\text{pH}(t)$  based on  $\text{TA}(t)$ ,  $\text{DIC}(t)$ , and water temperature using a standard Newton-Raphson iterative scheme (Stumm & Morgan, 2012).

## 2.2. Simplified Model

Equations 5–7 account for TA variation and its effect on the carbonate system. This may be desired to incorporate carbonate dynamics (e.g.,  $\text{HCO}_3^- + \text{H}^+ \leftrightarrow \text{CO}_2 + \text{H}_2\text{O}$ ) or to quantify biophysical processes affecting TA ( $g_{\text{TA}}(t)$ , e.g., calcification). However, in many cases of practical interest, TA can be approximated as constant ( $d\text{TA}/dt = 0$ ), simplifying calculations. In streams, such conditions typically require the following:

- Constant or baseflow discharge so that TA is in equilibrium with the lateral flow. Varying source activation under stormflow or hydrologically dynamic conditions likely induces temporal variation in TA.
- Negligible calcite dissolution and precipitation.

Additional considerations may be important depending on the model application that violate the constant TA assumption. For example, high catchment evapotranspiration may induce diel variation in alkalinity (an argument posited by Shangguan et al., 2021) with consequences for accurate modeling of the carbonate system. We note that direct  $\text{HCO}_3^-$  uptake by autotrophs under low  $\text{CO}_2$  conditions does not affect TA (Maberly & Gontero, 2017).

Additional model simplification is achieved when  $\text{CO}_3^{2-}$  is negligible (e.g.,  $\text{pH} < 8.5\text{--}9$ , depending on temperature), a common condition in many streams and rivers. Indeed, if one approximates total alkalinity in freshwater with the carbonate alkalinity  $\text{TA} \approx [\text{HCO}_3^-] + 2[\text{CO}_3^{2-}]$ , and  $\text{CO}_3^{2-} \approx 0$ , one has that  $\text{TA} \approx [\text{HCO}_3^-] \approx \text{constant}$ . Therefore

$$\frac{d\text{DIC}}{dt} = \frac{d\text{CO}_2}{dt} + \frac{d\text{HCO}_3^-}{dt} + \frac{d\text{CO}_3^{2-}}{dt} = \frac{d\text{CO}_2}{dt} \quad (8)$$

Under these assumptions and approximations, Equations 5–7 simplify to:

$$\frac{d\text{O}_2(t)}{dt} = \frac{\text{GPP}(t)}{z(t)} - \frac{\text{ER}(t)}{z(t)} + \frac{q_L(t)}{z(t)} [\text{O}_2^*(t) - \text{O}_2(t)] + K_{\text{O}_2}(t) [\text{O}_{2,\text{sat}}(t) - \text{O}_2(t)] \quad (9)$$

$$\begin{aligned} \frac{d\text{CO}_2(t)}{dt} = & -\frac{\text{GPP}(t)}{z(t)P_Q(t)} + R_Q(t) \frac{\text{ER}(t)}{z(t)} + \frac{q_L(t)}{z(t)} [\text{DIC}^*(t) - \text{HCO}_3^- - \text{CO}_2(t)] \\ & + K_{\text{CO}_2}(t) [\text{CO}_{2,\text{sat}}(t) - \text{CO}_2(t)] \end{aligned} \quad (10)$$

where  $\text{HCO}_3^-$  is the stream water bicarbonate concentration, which is constant under the above assumptions. The advantage of the assumption of negligible  $\text{CO}_3^{2-}$  is that the ODE system reduces to the most common state variables of interest:  $\text{O}_2$  and  $\text{CO}_2$ . Moreover, this model formulation eliminates the need to solve the carbonate equilibrium at each time step. This decreases model runtime, an important feature when performing Bayesian parameter estimation via Markov chain Monte Carlo.

### 2.3. Numerical Implementation

Equations 9 and 10 are integrated numerically with an Euler explicit scheme with a time step of 15 min. We checked the accuracy of the method by simulating the model with shorter time steps and comparing the results.

The temperature-dependent O<sub>2</sub> gas exchange rate is calculated from the parameter  $K_{600}$  (d<sup>-1</sup>), the gas exchange rate at the Schmidt number 600 (i.e., the Schmidt number of CO<sub>2</sub> at 20°C in freshwater):

$$K_{O_2}(t) = K_{600}(t) \left( \frac{600}{Sc_{CO_2}(T(t))} \right)^{0.5} \quad (11)$$

where  $Sc_{O_2}(T(t))$  is the function relating the Schmidt number for O<sub>2</sub>,  $Sc_{O_2}$ , to the water temperature  $T(t)$  (Wanninkhof et al., 2009).  $K_{CO_2}$  is computed using  $Sc_{CO_2}$ , the Schmidt number for CO<sub>2</sub>.

In the general application, the simplified model Equations 9 and 10 comprises 8 parameters, which need to be assigned or estimated (See Table 2). They are as follows: GPP, ER,  $q_L$ ,  $O_{2,L}$ ,  $DIC'_L$ ,  $K_{600}$ ,  $P_Q$ , and  $R_Q$ .  $DIC'_L(t) = DIC_L(t) - HCO_3^-$  is a combined parameter used in the estimation.

As customary in estimation of stream metabolism, (e.g., Hall & Hotchkiss, 2017) GPP is assumed to be a linear function of light  $L(t)$  according to

$$GPP(t) = GPP_d \frac{L(t)}{L_d} \quad (12)$$

where  $GPP_d$  and  $L_d$  are the mean of GPP and light values of day  $d$ , respectively. The units of  $L(t)$  are immaterial because of the normalization by  $L_d$ . The environmental forcings needed to run the model are  $L(t)$ ,  $T(t)$ , and  $z(t)$ .

Compared to typical applications of stream metabolism estimation based on diel O<sub>2</sub> observations, our model setup introduces two improvements. First, when observational error is assumed (see below), the simulation is run over several consecutive days while specifying only the initial conditions at the beginning of the simulation. Although setting separate initial conditions for each day is a valid approach—particularly when long time series are available or there are gaps in the data (see, e.g., Appling, Hall, et al., 2018)—using a single initial condition leverages the additional information of the continuity of O<sub>2</sub> and CO<sub>2</sub> time series across days. This ensures that there are no artificial step changes at midnight, reflecting the inherent smoothness of the processes. Second, parameters are assumed to change smoothly across days. Typical applications estimate separate parameter sets—such as mean daily ER, GPP, and  $K_{600}$ —for each day. While we adopt a similar approach, we also account for smooth day-to-day variations. Specifically, this is performed estimating the parameter values at noon of each day of the time series (e.g.,  $ER_d$ , and  $d = 1, \dots, N_d$ , where  $N_d$  is the number of days in the time series) and deriving the parameter value at each time (e.g.,  $ER(t)$ ) through linear interpolation of noon values. Thanks to these two improvements in the model setup the resulting simulated time series for O<sub>2</sub> and CO<sub>2</sub> are mathematical functions of class  $C^1$ : continuous and differentiable functions with continuous first derivatives.

### 2.4. Bayesian Parameter Estimation

Parameter estimation is performed in a Bayesian framework by sampling the parameter posterior distribution via a Markov chain Monte Carlo algorithm.

To test the model sensitivity to error structure, we employed two different error models: an observation error model and a process error model (for a detailed illustration of the two types of error in this context, the reader is referred to Appling, Hall, et al. (2018) and Song et al. (2016)). For the observation error model, the error,  $\epsilon_{obs}$ , between the simulated and observed concentrations of O<sub>2</sub> and CO<sub>2</sub>, is an independent, identically distributed normal variable with a mean of 0 and standard deviation  $\sigma$ ; for example,  $\epsilon_{O_2,obs}(t) = O_{2,mod}(t) - O_{2,obs}(t) = \mathcal{N}(0, \sigma_{O_2})$ . For the process error model, or the one-step-ahead model, the equations are used to evolve the system from an observation at time  $t$ , used as initial condition, to an estimate at time  $t + \Delta t$ , and the error, for example,  $\epsilon_{O_2,proc}$ , is calculated at time  $t + \Delta t$ :  $\epsilon_{O_2,proc}(t + \Delta t) = O_{2,mod}(t + \Delta t) - O_{2,obs}(t + \Delta t) = \mathcal{N}(0, \sigma_{O_2})$ .

We additionally tested model sensitivity to parameterization by comparing the simplified eight-parameter model with two other formulations. The first is the “classical” O<sub>2</sub> mass balance model with three parameters (GPP, ER, and  $K_{600}$ ). The second expands the classical model by introducing lateral discharge, and thus two additional parameters ( $q_L$  and O<sub>2,L</sub>). The additional model formulations are only conditioned on error in O<sub>2</sub>, and we compare GPP, ER, and  $K_{600}$  across the formulations. Finally, we assess the model sensitivity to the duration of the observed time series by repeating the estimation using half and one-quarter of the days in the data set.

We generally provide uninformative priors (uniform marginal distribution within a wide range) for parameters. When available and depending on the case study, informative priors are assigned to some parameters. In some other cases, priors are defined by hyperparameters, which follow hyperprior distributions in a hierarchical Bayesian framework. The detailed prior specification is reported after the three cases study are illustrated. Throughout the paper, results are reported as the median value and the 95-percentile range of the posterior distribution.

## 2.5. Case Studies

We searched the literature and data repositories for data sets containing jointly measured O<sub>2</sub> and CO<sub>2</sub> data, and we selected three suitable case studies, with varying hydrology and carbonate chemistries, to test the model (Table 1). We refer the readers to the original papers (Aho et al., 2021a; Kirk & Cohen, 2023) for additional information on each site, but note here that Union is a shallow, gravel-bed site on the Farmington River, Connecticut; Phelps is a small forested tributary to the Farmington River, located approximately 1 km upstream from a large reservoir; and SF2500 is a limestone-aquifer spring-fed river in Florida. From each site, we estimated model parameters using a time series with length varying from 8 to 13 days, chosen based on data quality and condition of baseflow. When light data were not available, we assumed light as proportional to the clear-sky solar radiation, which was calculated based on the hour of the day, day of year, latitude, and longitude (Thackston & Parker, 1971).

We estimated for each site lateral flow in order to provide informative priors for  $q_L$  and avoid model equifinality. To this end, we estimated  $q_L$  in one of two ways. (a) We differenced the discharge between the nearest ( $Q_1$ ) and the closest upstream United States Geological Survey (USGS) discharge gage ( $Q_0$ ). (b) We calculated the catchment area at the measurement site ( $A_1$ ) and the catchment area for a site located approximately 1,000 m upstream ( $A_0$ ). We then estimated the lateral discharge as  $Q_1(1 - A_0/A_1)$  assuming the two sites have the same specific discharge. To estimate  $q_L$  from these quantities, we divided this discharge by the distance between sites and by the average stream width as estimated from Google Earth.

Table 2 details the prior specifications for all parameters and case studies. After some initial trials, we note a strong equifinality at the Union and Phelps sites between the gas exchange rate  $K_{600}$  and metabolic fluxes GPP and ER. This is a known and common issue in the metabolism estimation (Appling, Hall et al., 2018) where equally good fitting of O<sub>2</sub> (and CO<sub>2</sub> in this case) can be achieved suitably increasing, or decreasing, both  $K_{600}$  and metabolic fluxes. To overcome this issue, we provide for these two sites an informative prior centered around the  $K_{600}$  estimated from hydraulic geometry (Raymond et al., 2012, Equation 1 in Table 2).

Two additional pairs of parameters that are difficult to independently estimate are  $q_L$  and DIC<sub>L</sub><sup>′</sup>, and  $q_L$  and O<sub>2,L</sub>, because their products appear in the model equations. For this reason, we use informative priors for  $q_L$  based on the estimates previously described. We further use a prior for O<sub>2,L</sub>, which assumes lateral flow is between oxygen saturation and depletion. No further assumption is instead done for DIC<sub>L</sub><sup>′</sup>.

For parameters DIC<sub>L</sub><sup>′</sup>,  $P_Q$ , and  $R_Q$ , we do not specify prior information, but we implemented a hierarchical Bayesian framework in order to allow day-to-day variability while maintaining interday consistency. We use a different parameter for each day but impose that they all follow the same prior distribution defined by hyperparameters, which are included in the estimation. Uninformative hyperpriors are assumed except for the parameters controlling the mean of the prior distribution of  $P_Q$  and  $R_Q$ . For these parameters, the hyperprior specifies that the mean of the prior distributions of  $P_Q$  and  $R_Q$  fluctuates around 1, the theoretical value for glucose synthesis and respiration.

Using the prior specification described above for site SF2500, we obtained quite large day-to-day variations in the parameters  $K_{600}$ ,  $R_Q$ , and DIC<sub>L</sub><sup>′</sup>, which seem unlikely given that the time series covers a period of almost constant baseflow discharge. To investigate if this is an identifiability issue, we estimated parameters in a simplified setup

**Table 1**  
Site Characteristics for Combined O<sub>2</sub> and CO<sub>2</sub> Data Sets Over the Model Fitting Period

Site	Period	$Q$	$q_L$	$z$	pCO <sub>2</sub>	pH	TA	Reference
Union	13-20/8/2019	4.2	0.09	0.4	928	7.6	0.57	(Aho et al., 2021b)
Phelps	10-20/8/2019	0.004	0.04	0.1	1,207	6.6	0.10	(Aho et al., 2021b)
SF2500	28/6–7/7/2019	33.7	0.57	1.7	6,520	6.7	2.0	(Kirk & Cohen, 2022)

*Note.*  $Q$  is mean stream discharge in [m<sup>3</sup> s<sup>-1</sup>] and  $q_L$  in [m d<sup>-1</sup>],  $z$  [m] is average cross-sectional water depth, mean pCO<sub>2</sub> in [ $\mu$  atm], mean DIC, and TA in [mol m<sup>-3</sup>] and [mEq m<sup>-3</sup>]. Lateral discharge was reported in the original data set for the site SF2500; it was estimated with the area method for the site Phelps, and with the discharge method for the site Union. For the latter, USGS gages 01188090, 01186000, and 01186500 were used. Average cross-sectional water depth was available in the original data set for the site SF2500, and was estimated via the hydraulic-stage discharge relationship for the other two sites (gage USGS 01187830 for Phelps and gage USGS 01188090 for Union).

and compared the results. Such setup (reported in Table 2) assumes a constant  $K_{600}$  and DIC<sub>L</sub>' for the 13-day long time series.

## 2.6. Code Details and Availability

Codes to run the model and reproduce the results are available in the public repository (<https://zenodo.org/doi/10.5281/zenodo.13193593>) in two programming languages: MatLab and Stan compiled in R. MatLab formulation uses the DREAM<sub>ZS</sub> (Vrugt et al., 2009) implementation of the Markov chain Monte Carlo algorithm. We ran 1 million simulations subdivided in three chains and used the last 25% of the runs to sample the posterior. Chain

**Table 2**  
Model Parameters, Associated Units, and Prior Distributions Assumed for the Three Case Studies

Parameter	Unit	Prior		
		Union	Phelps	SF2500
ER	mol m <sup>-2</sup> d <sup>-1</sup>	$\mathcal{U}(0, 3)$	$\mathcal{U}(0, 3)$	$\mathcal{U}(0, 3)$
GPP	mol m <sup>-2</sup> d <sup>-1</sup>	$\mathcal{U}(0, 3)$	$\mathcal{U}(0, 3)$	$\mathcal{U}(0, 3)$
$K_{600}$	d <sup>-1</sup>	$\mathcal{N}(13, 2)$	$\mathcal{N}(30, 5)$	$\mathcal{U}(0, 50)$
$q_L$	m d <sup>-1</sup>	$\mathcal{N}(0.1, 0.01)$	$\mathcal{N}(0.05, 0.01)$	$\mathcal{N}(0.6, 0.06)$
O <sub>2,L</sub>	mol m <sup>-3</sup>	$\mathcal{E}(0.05)$	$\mathcal{E}(0.05)$	$\mathcal{E}(0.05)$
DIC <sub>L</sub> '	mol m <sup>-3</sup>	$\mathcal{LN}(\mu_{\log(\text{DIC}'_L)}, \sigma_{\log(\text{DIC}'_L)})$	$\mathcal{LN}(\mu_{\log(\text{DIC}'_L)}, \sigma_{\log(\text{DIC}'_L)})$	$\mathcal{U}(0, 3)$
$P_Q$	–	$\mathcal{LN}(\mu_{\log(P_Q)}, \sigma_{\log(P_Q)})$	$\mathcal{LN}(\mu_{\log(P_Q)}, \sigma_{\log(P_Q)})$	$\mathcal{LN}(\mu_{\log(P_Q)}, \sigma_{\log(P_Q)})$
$R_Q$	–	$\mathcal{LN}(\mu_{\log(R_Q)}, \sigma_{\log(R_Q)})$	$\mathcal{LN}(\mu_{\log(R_Q)}, \sigma_{\log(R_Q)})$	$\mathcal{LN}(\mu_{\log(R_Q)}, \sigma_{\log(R_Q)})$
$\sigma_{\text{O}_2}$	mol m <sup>-3</sup>	$\mathcal{U}(0, 0.1)$	$\mathcal{U}(0, 0.1)$	$\mathcal{U}(0, 0.1)$
$\sigma_{\text{CO}_2}$	mol m <sup>-3</sup>	$\mathcal{U}(0, 0.1)$	$\mathcal{U}(0, 0.1)$	$\mathcal{U}(0, 0.1)$
Hyperparameter				
$\mu_{\log(\text{DIC}'_L)}$	log(mol m <sup>-3</sup> )	$\mathcal{U}(-5, 5)$	$\mathcal{U}(-5, 5)$	–
$\sigma_{\log(\text{DIC}'_L)}$	log(mol m <sup>-3</sup> )	$\mathcal{U}(0, 5)$	$\mathcal{U}(0, 5)$	–
$\mu_{\log(P_Q)}$	–	$\mathcal{N}(0, 0.2)$	$\mathcal{N}(0, 0.2)$	$\mathcal{N}(0, 0.2)$
$\sigma_{\log(P_Q)}$	–	$\mathcal{U}(0, 0.3)$	$\mathcal{U}(0, 0.3)$	$\mathcal{U}(0, 0.3)$
$\mu_{\log(R_Q)}$	–	$\mathcal{N}(0, 0.2)$	$\mathcal{N}(0, 0.2)$	$\mathcal{N}(0, 0.2)$
$\sigma_{\log(R_Q)}$	–	$\mathcal{U}(0, 0.3)$	$\mathcal{U}(0, 0.3)$	$\mathcal{U}(0, 0.3)$

*Note.*  $\mathcal{N}(\mu, \sigma)$  indicates a normal distribution with mean  $\mu$  and standard deviation  $\sigma$ ,  $\mathcal{LN}(\mu, \sigma)$  a lognormal distribution where  $\mu$  and  $\sigma$  are the mean and the standard deviation of the logarithm of the random variable, respectively,  $\mathcal{E}(\mu)$  an exponential distribution with a mean  $\mu$ , and  $\mathcal{U}(lb, ub)$  a uniform distribution in the range [ $lb, ub$ ]. Gray cells indicate where one parameter for each day of the time series has been estimated.

convergence was assessed via the Gelman and Rubin diagnostic. For the Union case study (8-day simulation at 15-min time step), 1 million simulations takes around 5 min in a standard laptop computer.

The STAN implementation uses a Hamiltonian Markov chain algorithm with adaptive no-U-turn sampling (*cmdstanr* package Stan Development Team (2023); Gabry et al. (2023)). For this approach, we used four parallel chains with 500 warm-up, or burn-in, samples and 2,000 posterior draws. For the Union case study, the sampling process takes about 80 s.

### 3. Results

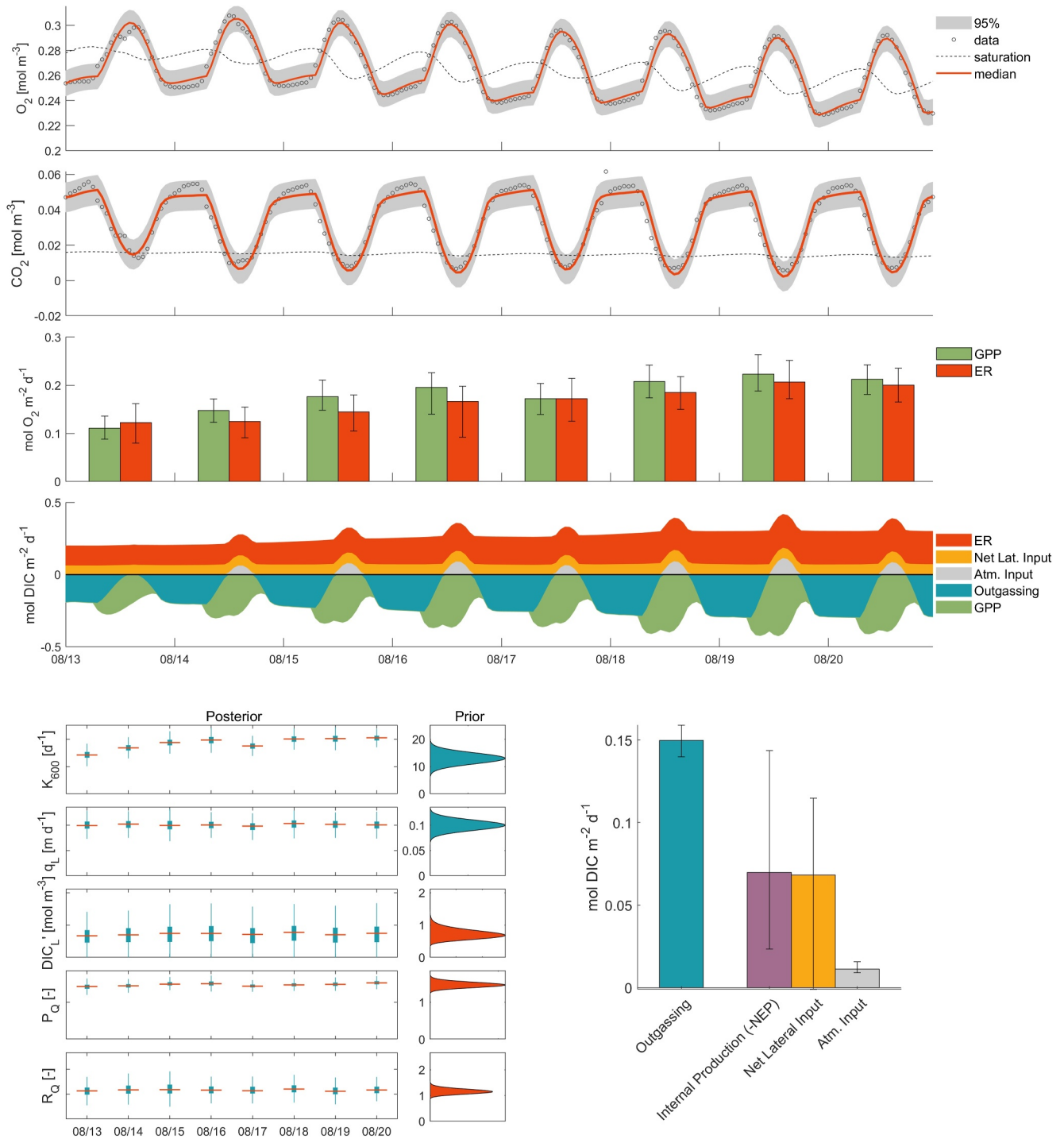
The simplified model (Equations 9 and 10) was able to replicate the main features of the O<sub>2</sub> and CO<sub>2</sub> time series for the three case studies. We report the root mean square error (RMSE) between observed and simulated concentrations to provide an easily interpretable measure of goodness of fit: median RMSE<sub>O<sub>2</sub></sub> = 0.0050, 0.0028, 0.0087 mol O<sub>2</sub> m<sup>-3</sup>, RMSE<sub>CO<sub>2</sub></sub> = 0.0041, 0.0020, and 0.0070 mol CO<sub>2</sub> m<sup>-3</sup> for sites Union, Phelps, and SF2500, respectively (Figures 2–4). When the same prior setup used for sites Union and Phelps was used for SF2500, there were high day-to-day parameter variations, in particular for  $K_{600}$ ,  $DIC'_L$ , and  $P_Q$  (Figure S4 in Supporting Information S1). We considered this unrealistic given that the time series covers a period of near-constant flow. The simplified setup produced instead more consistent results (Figure 4).

Both the observation error model and the process error model performed similarly and produced similar estimates (Figures S1–S3 in Supporting Information S1), although the process error model for SF2500 was less accurate than the observation error model. This is a well-known problem with process error models (Song et al., 2016), which fit changes in concentration accurately, but tend to perform poorly for actual concentration trajectories when continuous simulations are run. Process error models also tended to estimate lower metabolic rates than observation error models. Given the simpler model formulation and its generally better performance, we present the observation error model as the main results here.

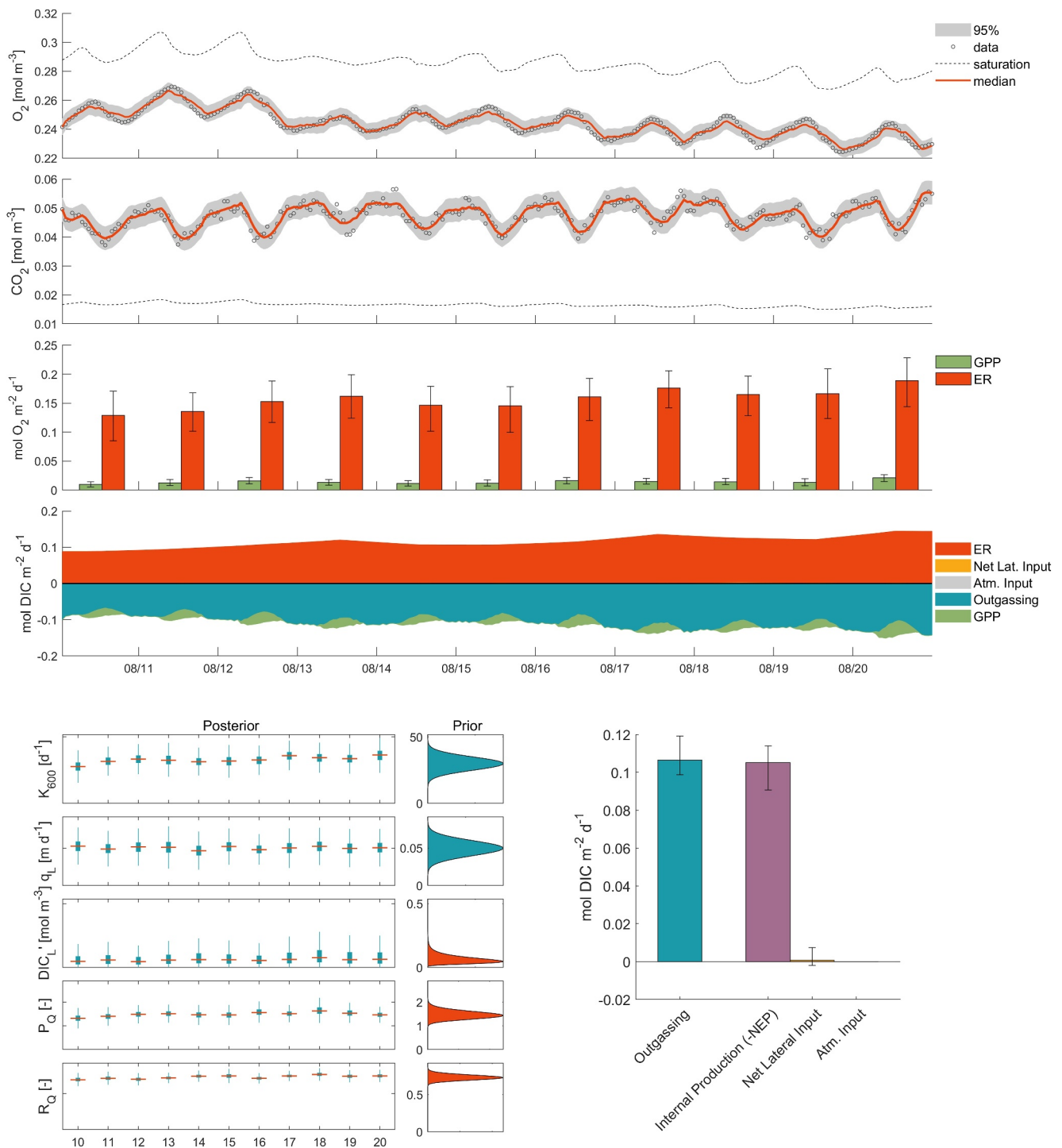
The duration of the data time series did not significantly affect the results (Figures S17–S19 in Supporting Information S1). Estimates obtained from shorter periods showed good agreement with those based on the full data set. As expected, uncertainty increased with shorter time series, particularly for parameters with a hierarchical structure (e.g.,  $DIC'_L$ ,  $R_Q$ , and  $P_Q$ ), where estimation relies on information aggregated across all days in the time series to determine the hyperparameters.

Model results allowed us to estimate the time series of input (ER, net lateral input, and atmospheric input) and output (GPP and outgassing) fluxes to the stream DIC balance, and from this, we partitioned the source of CO<sub>2</sub> outgassing for the whole simulated period (Figures 2–4). The lateral flow, when positive, constitutes an input to the DIC balance equal to  $Q_L DIC'_L$ . At the same time, positive lateral flow increases discharge and thus simultaneously increases net export of DIC:  $(Q - Q_u) DIC$  (the reader is reminded that under the model assumptions,  $DIC_u = DIC$ ). We term the net effect of these two processes as net lateral DIC contribution (Figures 2–4). Mathematically, this corresponds to  $q_L (DIC'_L - CO_2)$  when expressed as flux per unit of streambed area.

We observed some collinearity in the parameter posterior samples (Figures S11–S13 in Supporting Information S1). As anticipated, ER, GPP, and  $K_{600}$  were highly correlated ( $R^2 > 0.5$ ) for sites Union and Phelps. We also observed correlation between  $R_Q$  and  $DIC'_L$ , in particular for sites Union and SF2500, as both parameters control the input of DIC. From a mathematical viewpoint, their contributions to the DIC balance is different and thus potentially identifiable. The input  $ER \cdot R_Q$  is almost constant during the day (assuming low interday variation of parameters), while the term  $q_L (DIC'_L - CO_2)$  is modulated by the daily fluctuations of stream CO<sub>2</sub>. However, if  $DIC'_L \gg CO_2$ , like in the estimated values for Union and SF2500, the contribution related to  $DIC'_L$  is also fairly constant during the day, and the two input terms related to  $R_Q$  and  $DIC'_L$  become difficult to separate, hence the observed equifinality. For site SF2500, which was strongly affected by the lateral flow  $q_L$ , we observed collinearity between ER and O<sub>2,L</sub>. From a process viewpoint, this highlights how an increased respiration can be counterbalanced by an increased lateral oxygen concentration, leading to a similar O<sub>2</sub> trajectory. Finally, the model posterior sample for  $q_L$  tended to mirror the model prior, suggesting that CO<sub>2</sub> and O<sub>2</sub> data do not contain enough information to further constrain this parameter (Figures 2–4).



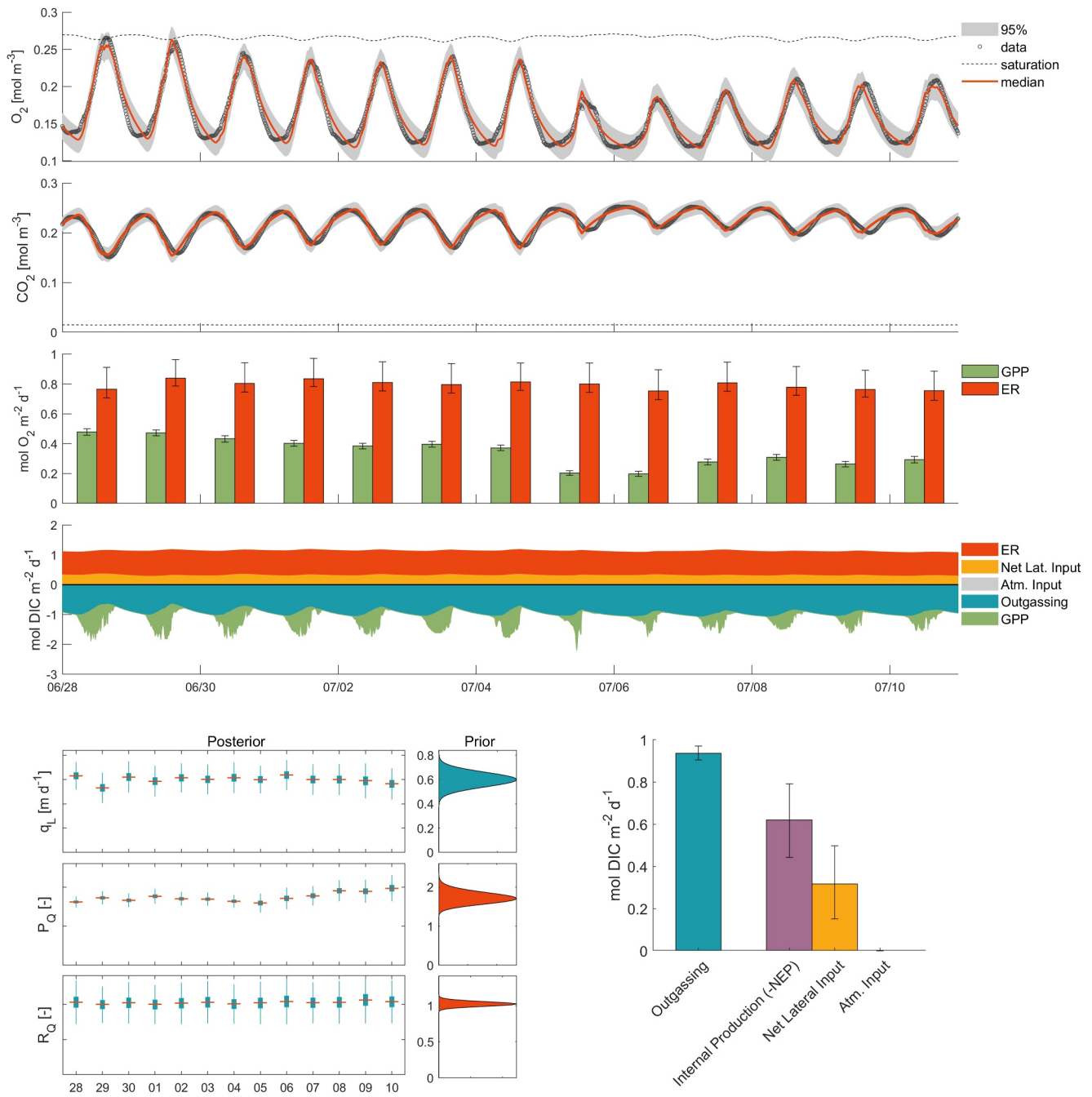
**Figure 2.** Model results for Farmington River at Unionville when observational error is assumed. From top to bottom: time series of  $O_2$ ,  $CO_2$ , mean daily GPP and ER, and median fluxes controlling the DIC balance (positive (negative) values represent input (output) fluxes). The bottom left panel shows parameter marginal posterior distributions and associated prior. Red color indicates distributions whose shape is defined by estimated hyperparameters. Prior and posterior distributions of the remaining parameters (i.e., those that are constant during the simulation) are presented in the accompanying Figure S5 of Supporting Information S1. Bottom right: mean values over the analyzed period of outgassing flux and its estimated sources. Throughout the figure, results are displayed as the median and the 95 percentile-range of the posterior density except for the time series of DIC fluxes where only the median values are shown to facilitate the visualization.



**Figure 3.** Results for the Phelps stream when observational error is assumed. Symbology as in Figure 2. Prior and posterior distributions of the remaining parameters (i.e., those that are constant during the simulation) are presented in the accompanying Figure S6 of Supporting Information S1.

### 3.1. Farmington River

The model was able to reproduce the finding of Aho et al. (2021a) that the Farmington River exhibited regular undersaturation of  $\text{CO}_2$  relative to the atmosphere leading to diurnal atmospheric inputs (Figure 2). This undersaturation was not due to particularly large rates of GPP (ca.  $0.1\text{--}0.2 \text{ mol O}_2 \text{ m}^{-2} \text{ d}^{-1}$ ).



**Figure 4.** Results for Santa Fe River when observational error is assumed. Symbology as in Figure 2. Prior and posterior distributions of the remaining parameters (i.e., those that are constant during the simulation) are presented in the accompanying Figure S7 of Supporting Information S1.

Model results allow computing NEP in terms of both  $O_2$  ( $NEP_{O_2} = GPP - ER$ ) and C ( $NEP_C = GPP/P_Q - ER \cdot R_Q$ ). The latter, with the opposite sign, represents the internal contribution to the DIC balance. For this site, we estimated a slightly autotrophic (i.e., positive NEP) system in terms of  $O_2$  ( $NEP_{O_2} = 0.014 (-0.001, 0.0036)$  mol  $O_2$   $m^{-2}$   $d^{-1}$ ). However, an estimated median  $P_Q$  of 1.47 and  $R_Q$  of 1.14 (Figure 2) translated this into a net heterotrophic metabolic system from a C perspective ( $NEP_C = -0.07 (-0.14, -0.02)$  mol C  $m^{-2}$   $d^{-1}$ ). Nearly equal DIC was estimated to be delivered to the river from lateral sources as from internal production (Figure 2), but there was large uncertainty in this proportioning. Estimates of GPP and ER were not particularly uncertain; however, their median values were close and thus their difference amplified the

uncertainty. In turn, due to the mass balance constraint, the lateral input was also uncertain and this translated to a wide posterior distribution for the parameter  $DIC'_L$ .

We note some difference between the GPP estimated here (0.18 (0.17, 0.19) mol O<sub>2</sub> m<sup>-2</sup> d<sup>-1</sup>) and in the original study (0.40 ± 0.3 mol O<sub>2</sub> m<sup>-2</sup> d<sup>-1</sup>). As the estimated  $K_{600}$  is similar, and our estimates based only on O<sub>2</sub> data are consistent with the estimates of the complete model (Figure S14 in Supporting Information S1), we attribute this difference to the possible different values used for water depth  $z$ . We reconstructed a discharge rating curve based on USGS channel data, while the original study could have used field knowledge to derive values of the water stage more representative of the whole upstream reach. However, we stress that in Equations 9 and 10 water depth represents a scale parameter that converts metabolic rates from a volumetric to an areal basis, and it does not affect the emission partitioning and the estimation of the other parameters.

### 3.2. Phelps Stream

Model results for the Phelps stream indicated nearly all of CO<sub>2</sub> degassed to the atmosphere derived from internal production (Figure 3). Estimates of GPP (0.014 (0.07, 0.023) mol O<sub>2</sub> m<sup>-2</sup> d<sup>-1</sup>) were in line with those of the original study (0.015 ± 0.015 mol O<sub>2</sub> m<sup>-2</sup> d<sup>-1</sup>). Also for this site, we observed no differences in GPP, ER, or  $K_{600}$  among model formulations (Figure S15 in Supporting Information S1). We estimated a credible difference between  $NEP_{O_2} = -0.14$  (-0.16, -0.11) mol O<sub>2</sub> m<sup>-2</sup> d<sup>-1</sup> and  $NEP_C = -0.10$  (-0.11, -0.09) mol C m<sup>-2</sup> d<sup>-1</sup>. Such difference is due to the estimated value of  $R_Q = 0.73$  (0.65, 0.80).  $P_Q$  was also different from 1 ( $P_Q = 1.48$  (1.14, 1.91)), but this did not affect  $NEP_C$  as GPP is almost negligible compared to ER.

### 3.3. Santa Fe River

Compared to the other sites, the SF2500 fit was less accurate and tended to lag O<sub>2</sub> and CO<sub>2</sub> observations during the afternoon/early evening, and exhibited steeper changes around daily minima and maxima than were observed (Figure 4). SF2500 exhibited consistent heterotrophy despite higher rates of GPP (0.34 (0.33, 0.35) mol O<sub>2</sub> m<sup>-2</sup> d<sup>-1</sup>, Figure 4). Our estimates of O<sub>2</sub>-based metabolism and  $K_{600}$  were closely aligned with those of the original authors (Kirk & Cohen, 2023). Estimated values of the photosynthetic quotient larger than unity ( $P_Q = 1.72$  (1.52, 2.07)) led to an increased heterotrophy in terms of C compared to O<sub>2</sub>:  $NEP_{O_2} = -0.45$  (-0.58, -0.40) mol O<sub>2</sub> m<sup>-2</sup> d<sup>-1</sup>;  $NEP_C = -0.62$  (-0.79, -0.44) mol C m<sup>-2</sup> d<sup>-1</sup>.

Differently from the other two case studies, we observed a sensitivity to model formulation, where “classical” O<sub>2</sub>-based estimates of both ER and  $K_{600}$  decreased by 15% with inclusion of  $q_L$  and by another 15% using the complete model (Figure S16 in Supporting Information S1). It was only with the full model, however, that  $K_{600}$  aligned with the original study's direct measurements. As we provided no informative prior of  $K_{600}$  at this site (Table 2), this highlights that the full model takes advantage of the increased information contained in CO<sub>2</sub> time series to more accurately constrain parameter posteriors.

After providing relatively tight priors for  $q_L$  based on the measurements of the original study, the model converged on a  $DIC'_L$  of 0.74 (0.46, 1.05) mol m<sup>-3</sup>.  $DIC'_L$  is a combined model parameter equal to  $DIC_L - HCO_3^-$ . To provide an approximate comparison with the original measurements of CO<sub>2,L</sub>, we can further assume that alkalinity in the stream is in equilibrium with that of the lateral flow, which may be valid during baseflow conditions in this spring-fed river. Under the hypothesis of minimal CO<sub>3</sub><sup>2-</sup>, this translates into  $HCO_3^- = HCO_{3,L}^-$ , and thus,  $DIC'_L = CO_{2,L} + HCO_{3,L}^- - HCO_3^- = CO_{2,L}$  (see Equation 10). Thus, the lower bound of our estimate of lateral CO<sub>2</sub> is in line with the measured in situ concentration (riparian flow 0.44 mol CO<sub>2</sub> m<sup>-3</sup>, and upland groundwater 0.30 mol CO<sub>2</sub> m<sup>-3</sup>) in the original study. Although this comparison is approximate, it is remarkable that the model was able to estimate the magnitude of the lateral CO<sub>2</sub> concentration solely based on in-stream measurements of O<sub>2</sub> and CO<sub>2</sub>. The greater C-based NEP contribution estimated here (the original study assumed  $P_Q = 1$ ) could help explaining the fraction of CO<sub>2</sub> outgassing that was not attributed to any source in Kirk and Cohen (2023).

## 4. Discussion

The model presented here represents an improved source tracking for stream and river CO<sub>2</sub> compared to previous inference-by-difference methods. Taking advantage of the combined information from CO<sub>2</sub> and O<sub>2</sub> time series,

we were able to partition CO<sub>2</sub> provenance between lateral inputs and internal production and to quantify the uncertainty in these fluxes. Model results indicated that internal production of CO<sub>2</sub> via NEP was, although to different degrees, the dominant source of stream CO<sub>2</sub> across three diverse streams during growing season conditions. While the limited set of case studies analyzed does not allow general conclusions, the tools proposed will hopefully facilitate a more systematic and quantitative assessment of CO<sub>2</sub> budget in flowing waters.

We designed the model used here to be flexible for a variety of approaches and needs, including the ability to account for the entire carbonate system, possible calcification effects, and observation or process error structure, and to be conditioned on O<sub>2</sub> and/or CO<sub>2</sub> observations. An additional, but unexplored, capacity of the model is to condition on pH data, which under most conditions is analogous to conditioning on CO<sub>2</sub>. Thus, for sites with only O<sub>2</sub> and pH measurements, the model may be equally useful, potentially opening its utility to a vast range of data sets, although estimations of TA are still required. One downside of this approach is that carbonate system calculations are required, thus increasing model runtime.

Lateral DIC inputs were one of the most uncertain fluxes in the results (Figures 2–4). We note that for the simplified model, the combined parameter, DIC'<sub>L</sub>, is approximately equal to CO<sub>2,L</sub> if stream water alkalinity is in equilibrium with lateral flow; otherwise HCO<sub>3,L</sub><sup>-</sup> and CO<sub>3,L</sub><sup>2-</sup> if present, can be an additional, but unquantified source of DIC that can rapidly be converted into CO<sub>2</sub> within the stream. While the simplified model accounts for the total contribution of lateral DIC, the role played by the different components of the water carbonate system cannot be unraveled in the current formulation. A more complete picture could be achieved using the full model (Equations 5–7), but at the cost of the need to estimate TA<sub>L</sub>. Alternatively, the need to solve the balance for TA may be avoided by using direct measurements, for example, with automated sensors (Shangguan et al., 2021), or with estimates derived from electrical conductivity.

The proposed model employs a “single station” approach, meaning it uses observed variables at a single cross-section to estimate representative quantities for the upstream reach, and the conditions for its applicability must be carefully evaluated. Typically, these conditions imply no major discontinuities or heterogeneities within a reach length of approximately  $L_u \sim 3V/K$ , where  $V$  is the average cross-sectional streamflow velocity, and  $K$  is the gas exchange rate. This distance represents the length a parcel of water travels before losing memory of perturbations and reequilibrating with the system (Reichert et al., 2009). Shangguan et al. (2024) analyze the applicability of single versus two-station approaches for CO<sub>2</sub>/DIC measurements in the presence of carbonate buffering. Specifically, carbonate buffering can slow the reequilibration time for CO<sub>2</sub> compared to O<sub>2</sub>, thereby increasing the length of the upstream reach that influences DIC dynamics at the sampled cross-section. This, in turn, can make the applicability of the single-station approach more challenging as O<sub>2</sub> and CO<sub>2</sub> integrate different stream lengths. The simplified model Equations 9 and 10 can provide further insights into this process. When the conditions for the applicability of the simplified model (negligible CO<sub>3</sub><sup>2-</sup> and constant alkalinity) are met, the equations for O<sub>2</sub> and CO<sub>2</sub> share the same structure. As a result, the reequilibration times for the two gases are similar, apart from the small difference between  $K_{O_2}$  and  $K_{CO_2}$ . This occurs because when alkalinity is constant and mostly composed by HCO<sub>3</sub><sup>-</sup>; variations in DIC are absorbed by variations in CO<sub>2</sub>:  $dDIC/dt = dCO_2/dt$ . When such conditions are not met,  $dDIC/dt \neq dCO_2/dt$ , and this results in a slowdown of the DIC reequilibration time compared to O<sub>2</sub>. As shown by Stets et al. (2017), this is more likely to occur at a high alkalinity, where CO<sub>3</sub><sup>2-</sup> can contribute significantly. Under these conditions, it is necessary to use the complete model (Equations 5 and 6) to account for carbonate equilibria reactions. In such cases, the conditions of the upstream reach and the DIC reequilibration time must be carefully evaluated to assess the applicability of a single-station approach. To verify these conditions, the development of a one-dimensional model of coupled DIC and CO<sub>2</sub> mass balances can be helpful. If the conditions are not satisfied, a two-station approach is recommended.

It should be emphasized that our approach assumes a lack of upstream discontinuities (e.g., tributaries or hydraulic jumps) not only for O<sub>2</sub> (as is typical in stream metabolism studies) but also for DIC and TA. The above discussion regarding the conditions for the applicability of the presented model could be relevant to the Phelps case study. Although the general expectation for a small stream is that CO<sub>2</sub> is predominately sourced from lateral inflow (Hotchkiss et al., 2015), we found that Phelps CO<sub>2</sub> outgassing could be almost completely sourced by internal production. This is possibly due to the peculiarity of the site and modeling time frames chosen: Phelps was characterized by low rates of lateral inflow during the summer, but winter might reveal contrasting behavior. It is also worth noting that while we estimate an upstream reach length  $L_u$  of approximately 150–200 m, the stream

crosses a wetland area 500–600 m upstream of the sampling station. As we were unable to verify in the field whether this wetland introduces discontinuities that persist to the sampling location, the results for this site should be interpreted with additional caution.

Model equifinality, while hindering parameter identification, sheds light on the physical processes that can lead to the same  $O_2$  and  $CO_2$  signals, highlighting important measurement gaps for future research efforts. Indeed, the equifinality between  $DIC_L$  and  $R_Q$ , highlights how, in many cases of practical interest,  $O_2$  and  $CO_2$  time series alone may not contain enough information to constrain these parameters. This equifinality animates the growing need for direct measurements of  $R_Q$  under varying environmental conditions to better inform prior distributions. In general, equifinality and uncertainty could be alleviated providing more informative prior distributions. Constrained estimates of lateral discharge and groundwater  $CO_2$  and DIC concentrations—and their variation with stream discharge—emerge as clear targets. Indeed, our efforts here highlight that  $q_L$  magnitude is a dominant control on model accuracy and thus accurate partitioning of  $CO_2$  provenance, invoking the need for regional and global estimates of lateral discharge. Moreover, hierarchical pooling of, for example,  $K_{600}$  versus discharge could be implemented to overcome equifinality (Appling, Hall, et al., 2018). Still, despite best efforts to reduce uncertainty, model equifinality may be intractable at certain scales under the current modeling framework, for example, due to the inherent correlation between  $CO_2$  and  $K_{600}$  (Rocher-Ros et al., 2019).

One of the important results of this work is the quantitative assessment of  $P_Q$  and  $R_Q$ , and their uncertainties. Despite only analyzing three sites, our results identify that  $P_Q$  and  $R_Q$  are rarely equal to 1. We note that while simple regressions of  $O_2$ – $CO_2$  can inform  $P_Q$  and  $R_Q$  (Aho et al., 2021a; Vachon et al., 2020), it remains difficult to parse  $P_Q$  from  $R_Q$  because (a) when estimating  $P_Q$  in daylight, the effect of both  $R_Q$  and  $P_Q$  are cooccurring, and (b) because  $R_Q$  in particular may be indistinguishable from signals deriving from variation in  $DIC_L$ , as evidenced by model equifinality. Finally, the simple physical differences in  $O_2$  and  $CO_2$  atmospheric gas exchange lead to regression slope magnitudes greater than 1 with increasingly greater magnitudes at higher temperatures, even when  $P_Q = R_Q = 1$  (see example in Figure S20 in Supporting Information S1). Thus, while  $O_2$ – $CO_2$  regression can represent a useful tool for an initial assessment, its results should be interpreted with caution and bearing in mind the potential confounding factors outlined above. By contrast, the proposed model accounts for all of these processes and their associated uncertainty, and indicates that  $P_Q$  is commonly at least 1.2 during the growing season, and that  $R_Q$  varies between less than 1 and more than 1 depending on site conditions. Considering only the improved estimate of  $P_Q$ , our model implies that previous estimates of internal  $CO_2$  production based on  $O_2$  signals may be underestimated. If  $R_Q$  is also greater than 1, then  $O_2$ -based estimates are biased even lower. Thus, our results suggest that internal  $CO_2$  production estimates based on  $O_2$  metabolic estimates need to properly account for the inherent uncertainty in  $P_Q$  and  $R_Q$ .

As the cost and reliability of in situ  $CO_2$  sensors improve, their deployment is expected to become as commonplace as  $O_2$  sensors are today. Consequently, we need models to evaluate the data streams they generate. Our model addresses this need by providing a flexible framework to estimate key parameters of the carbonate- $O_2$  system in flowing waters. Of particular novelty is its ability to simultaneously estimate  $P_Q$ ,  $R_Q$ , and lateral  $CO_2$  or DIC inflow—parameters that are typically assumed (as in the case of  $P_Q$  or  $R_Q$ ) or estimated by difference (as in the case of lateral DIC inflow). In the case studies presented,  $CO_2$  concentrations were measured using different methods: at site SF2500,  $CO_2$  was monitored with a sensor, whereas at Phelps and Union, it was measured using showerhead equilibrators, which are generally considered more accurate. Notably, we did not use prior information about  $CO_2$  measurement errors. For future applications involving sensor-based measurements, characterizing sensor error and incorporating this information as a prior could help prevent overconfidence in the data. Moreover, the Bayesian approach detailed here provides a straightforward way to quantify uncertainty in key parameters and the resulting carbon and oxygen fluxes. Potentially tractable improvements to our approach include the incorporation of state space modeling (i.e., accounting for both observation and process errors, see Appling, Hall, et al. (2018)), reductions in computational time, and mitigation of model equifinality as discussed above.

Finally, we hope that this model motivates efforts to deconvolve stream  $CO_2$  provenance at network scales. Initial efforts to do so may use a deterministic version of the model across reaches in synthetic stream networks (Carraro et al., 2020), where metabolism can be predicted across space and time (Diamond et al., 2021, 2023; Koenig et al., 2019; Mejia et al., 2019; Rodríguez-Castillo et al., 2019; Segatto et al., 2023). In such synthetic networks, it

will be necessary to develop network scaling functions for  $q_L$  (Kalbus et al., 2006; Leach et al., 2017). Additional needs will be network scaling functions for lateral DIC, TA, and CO<sub>2</sub> concentrations, which are currently poorly constrained. More care will also need to be given in the consideration of losing and gaining reaches and how best to incorporate spatial versus temporal changes in lateral inflow. The model presented here is an important step toward constraining terrestrial C leakage to freshwater systems, and to ultimately closing the linked terrestrial-aquatic carbon budget.

### Data Availability Statement

Data and codes to run the model and reproduce the results are available in the public repository: <https://zenodo.org/doi/10.5281/zenodo.13193593>.

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